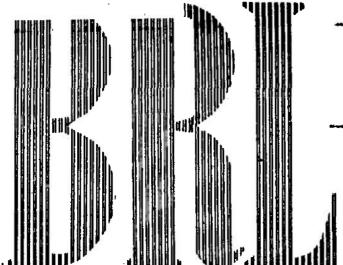
REFERENCE COPY

15 pg



REPORT NO. 1147 SEPTEMBER 1961

APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF
AN ALGEBRAIC IDENTITY

C. Masaitis

TECHNICAL LIBRARY
U S ARMY ORDRANCE
ABERDEEN PROVING GROUND, MD.
ORDBO-TL

[- - JUL 1996

REFERENCE COPY

DOES NOT CIRCULATE

Department of the Army Project No. 503-06-002
Ordnance Management Structure Code No. 5010.11.812
BALLISTIC RESEARCH LABORATORIES

ABERDEEN PROVING GROUND, MARYLAND

BALLISTIC RESEARCH LABORATORIES

REPORT NO. 1147

SEPTEMBER 1961

APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN ALGEBRAIC IDENTITY

C. Masaitis

Computing Laboratory

TECHNICAL LIBRARY
U. S. ARMY ORDNAHOE
ABERDEEN PROVING GROUND, MOORDEG TL

Department of the Army Project No. 503-06-002 Ordnance Management Structure Code No. 5010.11.812

ABERDEEN PROVING GROUND, MARYLAND

BALLISTIC RESEARCH LABORATORIES

REPORT NO. 1147

CMasaitis/sec Aberdeen Proving Ground, Md. September 1961

APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN ALGEBRAIC IDENTITY

ABSTRACT

An identity
$$\frac{x_1^n - x_2^n}{x_1^{-x_2}} = \sum_{i_1 + i_2 = n-1}^{i_1 + i_2 = n-1} x_1^{i_1} x_2^{i_2}$$

is here generalized to an arbitrary number of variables

 x_1, \ldots, x_m . The proof of the generalized identity

$$\sum_{i_1+i_2+\ldots+i_m=n-m+1}^{} x_1^{i_1} x_2^{i_2} \ldots x_m^{i_m} = \frac{x_1^n}{(x_1-x_2)(x_1-x_3)\ldots(x_1-x_m)} +$$

$$\frac{x_2^n}{(x_2-x_1)(x_2-x_3)\dots(x_2-x_m)} + \dots + \frac{x_m^n}{(x_m-x_1)(x_m-x_2)\dots(x_m-x_m-1)}$$

is obtained by computing transition probabilities in a Markov Chain.

Assume that $0 < x_i < 1$, i = 1, 2, ..., m, $x_i \neq x_j$ for $i \neq j$, and let $y_i = 1 - x_i$. Consider a Markov Chain with the matrix of transition probabilities:

$$\begin{pmatrix}
x_1 & y_1 & 0 & 0 & 0 & \dots & 0 \\
0 & x_2 & y_2 & 0 & 0 & \dots & 0 \\
0 & 0 & x_3 & y_3 & 0 & \dots & 0 \\
0 & 0 & 0 & 0 & 0 & \dots & x_m
\end{pmatrix}$$

If $n \ge r$ then a transition of the system from the state k to a state k+r in n steps can occur in any of the set of mutually exclusive way., each corresponding to a distinct sequence of integers t_1, t_2, \ldots, t_r , satisfying the inequalities $1 \le t_1 < t_2 < \ldots < t_r \le n$. In each of these cases the system passes from the state k+h to the state k+h+1 at the step number t_{h-1} . The probability of such an event is $y_k y_{k+1} \cdots y_{k+r-1} x_k x_{k+1} \cdots x_{k+r} \cdot Here the sum$ of the exponents of x_{k+1} 's is $t_1 - 1 + t_2 - t_1 - 1 + \ldots + n - t_r = n - r$. Therefore the probability $p_{kr}^{(n)}$ of transition in n steps from the state k to the state k+r is given by

(i)
$$p_{kr}^{(n)} = y_k y_{k+1} \dots y_{k+r-1} \sum_{\substack{i_1 + \dots + i_{r+1} = n-r \\ i_1 + \dots + i_{r+1} = n-r}} x_k^{i_1} x_{k+1}^{i_2} \dots x_{k+r}^{i_{r+1}}$$

On the other hand, $p_{kr}^{(n)}$ can be obtained by the method described in

[1]. Let D(s) be the determinant of the system

$$\xi_{j} = s(x_{j} \xi_{j} + y_{j} \xi_{j+1}), \quad j = 1, 2, ..., m - 1,$$
(2)
$$\xi_{m} = s x_{m} \xi_{m}.$$

and let $(\xi_1^{(t)}, \ldots, \xi_m^{(t)})$ be a non-trivial solution of (2) corresponding to the root s_t of D(s). Similarly, let $(\eta^{(t)}, \ldots, \eta_m^{(t)})$ be a non-trivial solution of

(3)
$$\eta_{1} = s x_{1} \eta_{1},$$

$$\eta_{i} = s(y_{i-1} \eta_{i-1} + x_{i} \eta_{i}),$$

corresponding to the same root s_t of D(s).

Let

(4)
$$c_t = \frac{1}{\sum_{\nu=1}^{m} \xi_{\nu}^{(t)} \eta_{\nu}^{(t)}}$$

and

(5)
$$\xi_{ji}^{(t)} = c_t \xi_j^{(t)} \eta_i^{(t)}$$

Then

(6)
$$p_{kr}^{(n)} = \frac{\zeta_{k k+r}}{s_1^n} + \frac{\zeta_{k k+r}}{s_2^n} + \dots + \frac{\zeta_{k k+r}}{s_m^n}$$

It follows from (2) that non-trivial solutions $(\xi_1^{(t)}, \ldots, \xi_m^{(t)})$ can be obtained for the following values of s: $s_1 = \frac{1}{x_m}$, $s_2 = \frac{1}{x_{m-1}}$,

..., $s_m = \frac{1}{x_1}$. By substituting $s_t = \frac{1}{x_{m-t+1}}$ in (2) the following expressions are obtained:

(7)
$$\xi_{m_i}^{(t)} = \xi_{m-1}^{(t)} = \dots = \xi_{m-t+2}^{(t)} = 0$$
,

(8)
$$\xi_{m-t+1}^{(t)} = 1$$
,

The substitution of (8) in the first equation of (2) yields

(9)
$$\xi_{m-t} = \frac{y_{m-t}}{x_{m-t+1} - x_{m-t}}$$

Suppose, by induction, that for $j \leq m - t$

(10)
$$\xi_{j}^{(t)} = \frac{m-t-j}{\left| \sum_{p=0}^{m-t+j} \frac{y_{j+p}}{x_{m-t+1}-x_{j+p}} \right|}$$

The substitution of (10) in (2) yields

$$\xi_{j-1}^{(t)} = \frac{\frac{m-t-j+1}{y_{j+p}}}{\frac{x_{m-t+1}-x_{j+p}}{y_{j+p}}}$$

Thus, in view of (9), (10) holds for every $j \leq m - t$.

Similarly, it follows from (3) that

$$\eta_{i}^{(t)} = 0 , \quad \text{for } i < m - t + 1 ;$$

$$\eta_{m-t+1}^{(t)} = 1 ;$$

$$\eta_{i}^{(t)} = \frac{i - m + t - 2}{\left| \frac{1}{m - t} \right|} \frac{y_{i-p-1}}{x_{m-t+1} - x_{j-p}} , \quad \text{for } i > m - t + 1$$

Therefore it follows from (4), (7), (8), and (11) that

(12)
$$c_t = 1$$
, $t = 1, 2, ..., m$.

Now (5), (7), (8), (10), and (12) yield

$$\zeta_{ji}^{(t)} = 0$$
, if $i \le m - t + 1$ or $j > m - t + 1$; $\zeta_{m-t+1 \ m-t+1}^{(t)} = 1$;

$$\zeta_{m-t+1 \ i}^{(t)} = \frac{\frac{m-t-i-2}{x_{m-t+1}-x_{i+p}}}{\frac{y_{i+p-1}}{x_{m-t+1}-x_{i+p}}}, \text{ if } i > m-t+1 ;$$

$$\zeta_{j \ m-t+1}^{(t)} = \frac{\frac{m-t-j}{x_{m-t+1}-x_{j+p}}}{\frac{y_{j+p}}{x_{m-t+1}-x_{j+p}}}, \text{ if } j < m-t+1$$

$$\zeta_{ij}^{(t)} = \frac{m-t-j}{p=0} \frac{y_{j+p}}{\frac{x_{m-t+1}-x_{j+p}}{x_{m-t+1}-x_{j+p}}} \frac{y_{i-q-1}}{\frac{y_{i-q-1}}{x_{m-t+1}-x_{j-q}}},$$

if i > m - t + 1 and j < m - t + 1.

Therefore by (6)

(13)
$$p_{1,m}^{(n)} = y_1 y_2 \dots y_{m-1} \left[\frac{x_m^n}{(x_m - x_{m-1})(x_m - x_{m-2}) \dots (x_m - x_1)} + \frac{x_{m-1}^n}{(x_{m-1} - x_m)(x_{m-1} - x_{m-2}) \dots (x_{m-1} - x_1)} + \dots + \frac{x_1^m}{(x_1 - x_m)(x_1 - x_{m-1}) \dots (x_1 - x_2)} \right]$$

Let
$$P_k(x_1, x_2, ...x_m) = \sum_{\substack{i_1+i_2+...+i_m=k}} x_1^{i_1} x_2^{i_2} ... x_m^{i_m}$$

Then by (1) and (13)

(14)
$$P_{n-m+1} (x_1, x_2, ..., x_m) = \frac{\begin{vmatrix} i & i & \cdots & i \\ x_1 & x_2 & \cdots & x_m \\ x_1^2 & x_2^2 & \cdots & x_m \\ x_1^{m-2} & x_2^{m-2} & \cdots & x_m^{m-2} \\ x_1^n & x_2^n & \cdots & x_m \end{vmatrix}}{V(x_1, x_2, ..., x_m)}$$

where V is a Vandermonde's determinant. Now the numerator N of the right hand side member of (14) is a polynomial in x_1, \ldots, x_m and it vanishes whenever $x_i = x_j$ for $i \neq j$. Since $V = (x_1 - x_2) \ldots (x_1 - x_m)$ ($x_2 - x_3$) ... $(x_{m-1} - x_m)$, it follows that N is divisible by V. Hence (14) is an equality of two polynomials which is valid for every x_i such that $0 < x_i < 1$. Hence (14) is an identity whenever N/V is defined, i.e. whenever $x_i \neq x_j$ for $i \neq j$. This identity can be written in the form

(15)
$$P_{n-m+1} = \frac{x_1^n}{(x_1 - x_2)(x_1 - x_3) \dots (x_1 - x_m)} + \frac{x_2^n}{(x_2 - x_1)(x_2 - x_3) \dots (x_2 - x_m)} + \frac{x_m^n}{(x_m - x_1)(x_m - x_2) \dots (x_m - x_{m-1})}.$$

For m = 2 (15) reduces to

$$x_1^{n-1} + x_1^{n-2}x_2 + x_1^{n-3}x_2^2 + \ldots + x_2^{n-1} = \frac{x_1^n}{x_1 - x_2} + \frac{x_2^n}{x_2 - x_1}$$

C. Masaitis

REFERENCE

W. Feller, An Introduction to Probability Theory and its
Applications, John Wiley and Sons, Inc., Chpt. 16.

DISTRIBUTION LIST

No. of Copies		No. of Copies	Organization
1	Chief of Ordnance ATTN: ORDTB - Bal Sec Department of the Army Washington 25, D.C.	1	Commander U.S. Naval Weapons Laboratory Dablgren, Virginia
1	Commanding Officer Diamond Ordnance Fuze Laboratories ATTN: Technical Information Office Branch O41		Chief of Staff U.S. Air Force The Pentagon Washington 25, D.C.
10	Washington 25, D.C. Commander Armed Services Technical Information Agency ATTN: TIPCR Arlington Hall Station	1	Commander Air Force Office of Scientific Research Building T-D Washington 25, D.C. Commanding Officer
10	Arlington 12, Virginia Commander British Army Staff British Defence Staff (W) ATTN: Reports Officer 3100 Massachusetts Avenue, N.W. Washington 8, D.C.	1	U.S. Army Research Office (Durham) Box CM, Duke Station Durham, North Carolina Army Research Office Arlington Hall Station Arlington, Virginia
l _t	Defence Research Member Canadian Joint Staff 2450 Massachusetts Avenue, N.W. Washington 8, D.C.	3	U.S. Atomic Energy Commission Los Alamos Scientific Laboratories P.O. Box 1663 Los Alamos, New Mexico
3	Chief, Bureau of Naval Weapons ATTN: DIS-33 Department of the Navy Washington 25, D.C.	1	Director National Bureau of Standards 232 Dynamometer Building Washington 25, D.C. Dr. C.V.L. Smith
2	Commander Naval Ordnance Laboratory White Oak, Silver Spring 19, Maryla		Data Reduction Division National Aeronautics and Space Administration 1520 H Street, N.W.
1	Commander U.S. Naval Ordnance Test Station ATTN: Technical Library China Lake, California		Washington 25, D.C.

AD Accession No. Ballistic Research Laboratories, APG APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN Identities. C. Massitis BRI Report No. 1147 September 1961 DA Proj No. 503-06-002, OMSC No. 5010.11.812 UNCLASSIFIED Report	AD Accession No. Ballistic Research Laboratories, APG ALGEBRAIC CONTINUATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN Identities, Markov C. Masaitis C. Masaitis BRL Report No. 1147 September 1961 DA Proj No. 503-06-002, CMSC No. 5010.11.812 UNCIASSIFIED Report
An identity $\frac{x_1^n \cdot x_2^n}{x_1^1 \cdot x_2^n} = \sum_{1_1+1_2=n-1}^{1_1} x_1^1 \cdot x_2^n$ is here generalized to an	An ide .ity $\frac{x_1^n - x_2^n}{x_1^n - x_2} = \sum_{1 + t_2 = n-1}^{-t_1} t_1^{-t_2} = \sum_{1 + t_2 = n-1}^{-t_1} t_2^{-t_2}$ is here generalized to an
arbitrary number of variables x_1, \ldots, x_m . The proof of the generalized identity $\sum_{1+12+\ldots+t_m=n-m+1}^{t_1} x_2^{1} \cdots x_m^{t_m} = \frac{x_1}{(x_1-x_2)(x_1-x_3)\cdots(x_1-x_m)}^{t_1}$	arbitrary number of variables x_1, \dots, x_m . The proof of the generalized identity $\sum_{\substack{1\\1+12}+\dots+1_m=n-m+1}}^{1} x_1^{1} x_2^{1} \dots x_m^{1} = \frac{x_1^n}{(x_1^n-x_2)(x_1^n-x_3^n)} + \frac{x_1^n}{(x_1^n-x_2)(x_1^n-x_3^n)} + \frac{x_1^n-x_2^n}{(x_1^n-x_2)(x_1^n-x_3^n)} + \frac{x_1^n-x_2^n}{(x_1^n-x_2^n)} + \frac{x_1^n-x_2^n}{(x_1^n-x_$
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	$ \begin{array}{cccccccccccccccccccccccccccccccccccc$
AD Ballistic Research Laboratories, APG APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN Identities, Markov ALGEBRAIC IDENTITY C. Massitis BRL Report No. 503-06-002, CMSC No. 5010.11.812 UNCLASSIFIED Report	AD Accession No. Ballistic Research Laboratories, APG APPLICATION OF A MARKOV CHAIN TO A GENERALIZATION OF AN ALGEBRAIC IDENTITY C. Massitis BRC. Report No. 1147 September 1961 DA Proj No. 503-06-002, OMSC No. 5010.11.812 UNCLASSIFIED Report
An identity $\frac{x_1^{-}x_2^{n}}{x_1^{-}x_2^{n}} = \sum_{1_1+1_2=n-1}^{1_1} x_1^{-1_2} x_2^{-1_2}$ is here generalized to an	An identity $\frac{x_1^n - x_2^n}{x_1^n - x_2} = \sum_{1 = 1/2}^{\infty} x_1^{-1} x_2^{-1}$ is here generalized to an
arbitrary number of variables x_1, \ldots, x_n . The proof of the generalized identity $\sum_{\substack{1\\1+t_2+\ldots+t_m=n-m+1}} t_1 x_2 \dots x_m = \frac{x_1}{(x_1-x_2)(x_1-x_3)\dots(x_1-x_m)} +$	arbitrary number of variables x_1, \dots, x_n . The proof of the generalized identity $\sum_{1_1+1_2+\dots+1_m=n-m+1}^{t_1} x_2^{1_1} x_2^{1_2} \dots x_m^{n-n} = (x_1-x_2)(x_1-x_2) \dots (x_1-x_m)$
x_{2}^{D} $(x_{2}-x_{1})(x_{2}-x_{3})\dots(x_{2}-x_{m}) + \dots + (x_{m}-x_{1})(x_{m}-x_{2})\dots(x_{m}-x_{m-1})$ is obtained by computing transition probabilities in a Markov Chain.	x_2 $(x_2-x_1)(x_2-x_3)\dots(x_2-x_m)$ $+\dots+(x_m-x_1)(x_m-x_2)\dots(x_m-x_{m-1})$ is obtained by computing transition probabilities in a Markov Chain.